## Homework 10 Solution

April 30th

## Exercises from the book

**Exercise 2.32** Since the density function f(x) is odd we see that  $\mathbb{E}[X] = 0$ .

$$\operatorname{Var}(X) = \mathbb{E}[X^2] = \int_{-\infty}^{\infty} x^2 f(x) \, dx = \int_0^{\infty} x^2 e^{-x} \, dx$$

since  $x^2 f(x)$  is even. Integrating by parts twice gives Var(X) = 2. Alternatively, we can also notice that this last integral is the second moment of an exponential distribution of parameter 1; for  $Y \sim \mathcal{E}xp(1)$ ,

$$\operatorname{Var}(X) = \int_0^\infty x^2 e^{-x} \, dx = \mathbb{E}[Y^2] = \operatorname{Var}(Y) + \mathbb{E}[Y]^2 = 1 + 1^2 = 2.$$

**Exercise 2.34** Denote by X the random location where you break the stick, and by L = L(X) the length of the longest of the two resulting segments. Then

$$L(X) = \begin{cases} 1 - X, & X \le \frac{1}{2}, \\ X, & X > \frac{1}{2}. \end{cases}$$

We have

$$\mathbb{E}[L] = \int_0^{\frac{1}{2}} (1-x)dx + \int_{\frac{1}{2}}^1 x dx = -\frac{(1-x)^2}{2} \bigg|_0^{1/2} + \frac{x^2}{2} \bigg|_{1/2}^1 = \frac{3}{4}.$$

**Exercise 2.36** Suppose that the lifetime T of a light bulb is exponentially distributed, so  $T \sim \mathcal{E}xp(\lambda)$ . The probability that that a bult will last more that a year is  $\mathbb{P}(T > 1) = e^{-\lambda}$  so we must have

$$0.8 \approx e^{-\lambda}.$$

The probability that a light bulb will last more than two years must be  $e^{-2\lambda}$  so we must have

$$0.3 \approx e^{-2\lambda} = (e^{-\lambda})^2 \approx (0.8)^2$$
.

This is clearly not the case so the lifetime is not exponentially distributed.

**Exercise 2.37** The easy way is to not try and compute the cumulative distribution function of X nor Y. Since X cannot be negative, Y cannot either, so  $p_Y(y) = 0$  for  $y \leq 0$ . We have

$$F_Y(y) = \mathbb{P}(\lambda X \le y) = \mathbb{P}(X \le y/\lambda) = F_X(y/\lambda)$$

Taking derivatives for y > 0, we see that

$$p_Y(y) = F'_Y(y) = \frac{\mathrm{d}}{\mathrm{d}y} \Big( F_X(y/\lambda) \Big) = p_X \Big( \frac{y}{\lambda} \Big) \cdot \frac{1}{\lambda} = \mathrm{e}^{-y}.$$

Thus,  $Y = \lambda X$  has the same cumulative distribution function as  $\mathcal{E}xp(1)$ .

If we wanted to follow the approach of the exercise, we would find  $F_X(x) = 0$  for  $x \le 0$ , and for x > 0

$$F_X(x) = \int_{-\infty}^x p_X(u) \mathrm{d}u = \int_0^x \lambda \mathrm{e}^{-\lambda u} \mathrm{d}u = 1 - \mathrm{e}^{-\lambda x}.$$

Because of the above, this would mean  $F_Y(y) = F_X(y/\lambda) = 1 - \exp(-y)$ , the derivative of which being what we expected indeed.

**Exercise 2.45** Let  $Y = X^2$ . Set  $F_Y(y) := \mathbb{P}(Y \leq y)$ . Because  $0 \leq Y \leq 1$ , we know that  $p_Y(y) = 0$  for  $y \notin (0, 1)$ . For 0 < y < 1, we have

$$F_Y(y) = \mathbb{P}(X^2 \le y) = \mathbb{P}(-\sqrt{y} \le X \le \sqrt{y}) = F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

for  $F_X$  the cumulative distribution function of X. The probability density function of Y is  $F'_Y(y)$ , so

$$F'_{Y}(y) = \frac{1}{2\sqrt{y}} \left( F'_{X}(\sqrt{y}) + F'_{X}(-\sqrt{y}) \right)$$
$$= \frac{1}{2\sqrt{y}} \left( \frac{1}{1 - (-1)} + \frac{1}{1 - (-1)} \right)$$
$$= \frac{1}{2\sqrt{y}}.$$

In particular, Y is in fact a beta distribution with parameters  $(\frac{1}{2}, 1)$ .

If we wanted to follow the approach of the exercise, we would find  $F_X(x) = 0$  for  $x \leq -1$ ,  $F_X(x) = 1$  for  $x \geq 1$ , and for 0 < x < 1 we have

$$F_X(x) = \int_{-\infty}^x p_X(u) du = \int_{-1}^x \frac{1}{2} du = \frac{x+1}{2}.$$

Because of the above, this would mean  $F_Y(y) = F_X(\sqrt{y}) - F_X(-\sqrt{y}) = \sqrt{y}$ , the derivative of which being what we expected indeed.

## Exercise 1

Write Y for the highest bid of the other contestants, so  $Y \sim Unif([70, 130])$ .

1. If the bid x is at most 70, the probability is zero, and it is one if x is at least 130. If 70 < x < 130, the probability of winning is

$$\mathbb{P}(\text{winning}) = \mathbb{P}(Y < x) = \int_{-\infty}^{x} p_Y(y) dy = \int_{70}^{x} \frac{1}{130 - 70} dy = \frac{x - 70}{60}.$$

All in all, we get

$$\mathbb{P}(\text{winning}) = \begin{cases} 0 & \text{for } x \le 70, \\ \frac{x - 70}{60} & \text{for } 70 < x < 130, \\ 1 & \text{for } x \ge 130. \end{cases}$$

2. Write X for the gain. For x fixed, X can take at most two values: 0 or 100 - x. The expectation is then

$$\mathbb{E}[X] = 0 \cdot \mathbb{P}(\text{losing}) + (100 - x) \cdot \mathbb{P}(\text{winning}) = \begin{cases} 0 & \text{for } x \le 70\\ \frac{(x - 70)(100 - x)}{60} & \text{for } 70 < x < 130\\ 100 - x & \text{for } x \ge 130. \end{cases}$$

3. For  $x \ge 130$ , the gain is negative. We have to study (x - 70)(100 - x). For instance, we can see that

$$(x - 70)(100 - x) = 225 - (x - 85)^2,$$

so that (x - 70)(100 - x) is at most 225, and it the maximum is reached at x = 85. Since 100 - 225/60 = 385/4 is positive, it is the maximal possible expectation, and the most profitable bid in average is 85.

## Exercise 2

1. A function p is the probability density function of a continuous variable if and only if p is nonnegative and its integral is equal to one. So we need a and b to be nonnegative, and

$$1 = \int_{-\infty}^{+\infty} p_X(x) dx = \int_{-1}^{1} a dx + \int_{0}^{1} b dx = a + b.$$

2. Let us compute the first two moments of X.

$$\mathbb{E}[X] = \int_{-1}^{0} ax \, dx + \int_{0}^{1} bx \, dx = \frac{ax^{2}}{2} \Big|_{-1}^{0} + \frac{bx^{2}}{2} \Big|_{0}^{1} = \frac{b-a}{2}$$
$$\mathbb{E}[X^{2}] = \int_{-1}^{0} ax^{2} \, dx + \int_{0}^{1} bx^{2} \, dx = \frac{ax^{3}}{3} \Big|_{-1}^{0} + \frac{bx^{3}}{3} \Big|_{0}^{1} = \frac{a+b}{3}$$

Hence, the variance of X is

$$\operatorname{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{4(a+b) - 3(b-a)^2}{12}.$$

Remember that a + b = 1, so the variance is minimal when |b - a| is maximal, which means that (a, b) equals either (0, 1) or (1, 0).

**Exercise 3** See Exercise 2.37 and Exercise 2.45 above.